

## 2- Gamma distribution

**Def :** If  $\alpha > 0$ , we define the gamma function  $\Gamma_\alpha = \int_0^\infty y^{\alpha-1} e^{-y} dy$

Properties of gamma function :

- 1-  $\Gamma_{(a+1)} = \alpha \Gamma_\alpha$
- 2- If  $\alpha$  is positive integer then  $\Gamma_{(\alpha+1)} = \alpha !$
- 3-  $\Gamma_\alpha = 2 \int_0^\infty x^{2\alpha-1} e^{-x^2} dx$
- 4-  $\Gamma_{\frac{1}{2}} = \sqrt{\pi}$

**Note :**

- 1- If  $\alpha = 1 \Rightarrow \Gamma(\alpha) = \int_0^\infty y e^{\alpha-1} dy = 1$
- 2- If  $\alpha > 1 \Rightarrow \pi(x) = (a-r) \int_0^\infty y^{\alpha-2} e^{-y} dy = (\alpha-1) \pi^{(\alpha-1)} = \pi(\alpha) = (\alpha-1)!$

**Def :** the continuous r.v  $x$  is said to have a gamma distribution with parameters  $\alpha$ ,  $\beta > 0$  denoted as  $x \sim G(\alpha, \beta)$  if the p. d. f of  $x$  is  $f(x) = \begin{cases} \frac{1}{\Gamma_\alpha \beta^\alpha} x^{\alpha-1} e^{-x/\beta}, & x > 0 \\ 0 & \text{o.w} \end{cases}$

**Properties of gamma distribution: -**

- 1- The m. g. f of the distribution is  $M_x(t) = (1 - \beta t^{-\alpha})^{-1}$ ,  $t < \frac{1}{\beta}$

**Proof :**  $M_x(t) = E(e^{tx}) = \int_0^\infty e^{tx} f(x) dx = \int_0^\infty e^{tx} \frac{1}{\Gamma_\alpha \beta^\alpha} x^{\alpha-1} e^{-\frac{x}{\beta}} dx$

$$= \frac{1}{\Gamma_\alpha \beta^\alpha} \int_0^\infty x^{\alpha-1} e^{tx - \frac{x}{\beta}} dx = \frac{1}{\Gamma_\alpha \beta^\alpha} \int_0^\infty x^{\alpha-1} e^{\frac{\beta tx - x}{\beta}} dx$$

$$= \frac{1}{\Gamma_\alpha \beta^\alpha} \int_0^\infty x^{\alpha-1} e^{-x \left( \frac{1-\beta t}{\beta} \right)} dx$$

Putting

$$y = x \left( \frac{1-\beta t}{\beta} \right) \Rightarrow x = \frac{\beta}{1-\beta t} y, dx = \frac{\beta}{1-\beta t} dy$$

$$M_x(t) = \frac{1}{\Gamma_\alpha \beta^\alpha} \int_0^\infty \left[ \frac{\beta y}{1-\beta t} \right]^{\alpha-1} e^{-y} \frac{\beta}{1-\beta t} dy$$

$$= \frac{1}{\Gamma_\alpha \beta^\alpha} \frac{\beta^\alpha}{(1-\beta t)^\alpha} \int_0^\infty e^{-y} y^{\alpha-1} dy$$

$$= \frac{1}{\Gamma_\alpha \beta^\alpha} \left( \frac{\beta}{1-\beta t} \right)^\alpha \Gamma_\alpha = \frac{1}{(1-\beta t)^\alpha} = (1 - \beta t)^{-\alpha}$$

2-  $M_x = E(x) = \alpha \beta$

3-  $\delta x^2 = \text{ver}(x) \cdot \alpha \beta^2$                       prove it

4- the  $k^{\text{th}}$  moment about origin is  $E(x^k) = \frac{\beta^k \Gamma_{\alpha+k}}{\Gamma_{\alpha}}$ ,  $k = 1, 2, 3, \dots$

**proof :**  $E(x^k) = \int_0^{\infty} x^k f(x) dx = \int_0^{\infty} x^k \frac{1}{\Gamma_{\alpha} \beta^{\alpha}} x^{\alpha-1} e^{-x/\beta} dx$

$$= \frac{1}{\Gamma_{\alpha} \beta^{\alpha}} \int_0^{\infty} x^{k+\alpha-1} e^{-x/\beta} dx$$

Let  $y = \frac{x}{\beta} \Rightarrow x = \beta y \Rightarrow dx = \beta dy$

$$E(x^k) = \frac{1}{\Gamma_{\alpha} \beta^{\alpha}} \int_0^{\infty} (\beta y)^{\alpha+k-1} e^{-y} (\beta dy)$$

$$= \frac{1}{\Gamma_{\alpha} \beta^{\alpha}} \beta^{\alpha+k} \int_0^{\infty} y^{\alpha+k-1} e^{-y} dy \qquad \int_0^{\infty} y^{\alpha-1} e^{-y} dy = \Gamma_{\alpha}$$

$$= \frac{1}{\Gamma_{\alpha} \beta^{\alpha}} \beta^{\alpha} \beta^k \Gamma_{(\alpha+k)} = \frac{\beta^k \Gamma_{\alpha+k}}{\Gamma_{\alpha}}, \quad k = 1, 2, \dots$$

**Ex:** use the formula of  $(x^k)$ , to find  $\mu_x, \delta x^2$

**Solution :** putting  $k = 1$  then

$$E(x) = M_x \frac{\beta \Gamma_{\alpha+1}}{\Gamma_{\alpha}} = \frac{\beta \alpha \Gamma_{\alpha}}{\Gamma_{\alpha}} = \alpha \beta$$

Putting  $k = 2$ , we get

$$E(x^2) = \frac{\beta^2 \Gamma_{(\alpha+2)}}{\Gamma_{\alpha}} = \frac{\beta^2 (\alpha+1) \Gamma_{(\alpha+1)}}{\Gamma_{\alpha}}$$

$$= \frac{\beta^2 (\alpha+1) \alpha \Gamma_{\alpha}}{\Gamma_{\alpha}} = \beta^2 \alpha (\alpha + 1)$$

$$\delta x^2 = \text{var}(x) = E(x^2) - [E(x)]^2$$

$$= \beta^2 (\alpha + 1) \alpha - \beta^2 \alpha^2$$

$$= \beta^2 \alpha^2 + \beta^2 \alpha - \beta^2 \alpha^2 = \alpha \beta^2$$

### 3- the chi square distribution

The chi square distribute on is as pecial case of gamma distribution which  $\alpha = \frac{r}{2}$  and  $B=2$  where ris positive integer . Hence the p.d.f of the r. v. x is

$$f(x) = \begin{cases} \frac{1}{\Gamma_{\frac{r}{2}} 2^{r/2}} x^{\frac{r}{2}-1} e^{-\frac{x}{2}}, & x > 0 \\ 0 & \text{o.w} \end{cases}$$

and we write  $x \sim \chi^2(r)$  where r is the number of degrees freedom representing the parameter of the distribution .

#### properties :-

the properties of chi square dist are the same of properties of gamma dist when  $\alpha = \frac{r}{2}$  ,  $\beta = 2$  that is :-

- 1)  $M_x(t) = (1 - 2t)^{-r/2}$
- 2)  $M_x = E(x) = \frac{r}{2} \cdot 2 = r$
- 3)  $\text{Var}(x) = \delta x^2 = \left(\frac{r}{2}\right) 2^2 = 2r$
- 4)  $E(x^k) = \frac{2^k \Gamma_{\left(\frac{r}{2}+k\right)}}{\Gamma_{\frac{r}{2}}}$  ,  $k = 1, 2, \dots$

### 4-Beta distribution

**Def :-** If  $\alpha > 0$  ,  $\beta > 0$  , we define the Beta function as

$$\beta(\alpha, \beta) = \int_0^1 x^{\alpha-1} (1-x)^{\beta-1} dx$$

It can be show that  $\beta(\alpha, \beta) = \frac{\Gamma_{\alpha} \Gamma_{\beta}}{\Gamma_{\alpha+\beta}}$  ,  $\alpha, \beta > 0$

**Def :-** the continuous r. v. x is said to have a Beta distribution denoted as  $x \sim \beta(\alpha$

,  $\beta)$  If the p. d. f of x is  $f(x) = \begin{cases} \frac{\Gamma_{\alpha+\beta}}{\Gamma_{\alpha} \Gamma_{\beta}} x^{\alpha-1} (1-x)^{\beta-1}, & 0 < x < 1 \\ 0 & \text{o.w} \end{cases}$

#### Properties :-

- 1) The  $k^{th}$  moment a bout origin is  $E(x^k) = \frac{\Gamma_{(k+\alpha)} \Gamma_{(\alpha+\beta)}}{\Gamma_{\alpha} \Gamma_{(k+\alpha+\beta)}}$  ,  $|k = 1, 2, \dots$

$$\begin{aligned}
\text{Proof: } E(x^k) &= \int_0^1 x^k f(x) dx = \frac{\Gamma_{\alpha+B}}{\Gamma_{\alpha}\Gamma_B} \int_0^1 x^{\alpha-1-k} (1-x)^{\beta-1} dx \\
&= \frac{\Gamma_{\alpha+\beta}}{\Gamma_{\alpha}\Gamma_{\beta}} \frac{\Gamma_{\alpha+k}\Gamma_{\beta}}{\Gamma_{\alpha+k+\beta}} \left[ \beta(\alpha, \beta) = \frac{\Gamma_{\alpha}\Gamma_{\beta}}{\Gamma_{\alpha+\beta}} \right] \\
&= \frac{\Gamma_{\alpha+B}\Gamma_{\alpha+k}}{\Gamma_{\alpha}\Gamma_{k+\alpha+\beta}}
\end{aligned}$$

2) From the above formula , the mean and variance of the distribution can be derive as follows :- putting k=1 we obtain

$$E(x) = M_x = \frac{\Gamma_{\alpha+1}\Gamma_{\alpha+\beta}}{\Gamma_{\alpha}\Gamma_{\alpha+\beta+1}} = \frac{\alpha\Gamma_{\alpha}\Gamma_{\alpha+\beta}}{\Gamma_{\alpha}(\alpha+\beta)\Gamma_{\alpha+\beta}} \text{ [since } \Gamma_{\alpha+1} = \alpha \Gamma_{\alpha}]$$

$$M_x = \frac{\alpha}{\alpha+\beta}$$

$$\text{Putting } k=2 \text{ we get } E(x^2) = \frac{\Gamma_{\alpha+2}\Gamma_{\alpha+\beta}}{\Gamma_{\alpha}\Gamma_{\alpha+\beta+2}} = \frac{(\alpha+1)\alpha\Gamma_{\alpha}\Gamma_{\alpha+\beta}}{\Gamma_{\alpha}(\alpha+\beta+1)(\alpha+\beta)\Gamma_{\alpha+\beta}}$$

$$= \frac{\alpha(\alpha+1)}{(\alpha+\beta+1)(\alpha+\beta)}$$

$$\text{Var}(x) = \int_x^2 = E(x^2) - [E(x)]^2$$

$$= \frac{\alpha(\alpha+1)}{(\alpha+\beta+1)(\alpha+\beta)} - \frac{\alpha^2}{(\alpha+\beta)^2} = \frac{\alpha(\alpha+1)(\alpha+\beta) - \alpha^2(\alpha+\beta+1)}{(\alpha+\beta)^2(\alpha+\beta+1)}$$

$$= \frac{\alpha^3 + \alpha^2\beta + \alpha^2 + \alpha\beta - \alpha^3 - \alpha^2\beta - \alpha^2}{(\alpha+\beta)^2(\alpha+\beta+1)} = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$$

$$\therefore \int_x^2 = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$$

## 5- The normal distribution

A continuous r.v x is said to have normal distribution with parameters M ,  $\delta^2$  denote , as  $x \sim N(M, \delta^2)$  if the p. d. f of x is  $f(x) = \frac{1}{\sqrt{2\pi\delta^2}} e^{-\frac{1}{2}\left(\frac{x-M}{\delta}\right)^2}$   $-\infty < x < \infty$

### Proportion :-

1) The M. g. f of normal dist. Is  $M_x(t) = e^{Mt - \frac{\delta^2 t^2}{2}}$

$$\text{Proof :- } M_x(t) = E(e^{tx}) = \frac{1}{\sqrt{2\pi\delta^2}} \int_{-\infty}^{\infty} e^{+x} e^{-\frac{1}{2}\left(\frac{x-M}{\delta}\right)^2} dx$$